A systematization of the Unscented Kalman Filter theory
Henrique M. T. Menegaz, João Y. Ishihara, Geoavany A. Borges, Alessandro N. Vargas

Abstract—In this paper, we propose a systematization of the (discrete-time) Unscented Kalman Filter (UKF) theory. We gather all available UKF variants in the literature, present corrections to theoretical inconsistencies, and provide a tool for the construction of new UKF’s in a consistent way. This systematization is done, mainly, by revisiting the concepts of Sigma-Representation, Unscented Transformation (UT), Scaled Unscented Transformation (SUT), UKF, and Square-Root Unscented Kalman Filter (SRUKF). Inconsistencies are related to 1) matching the order of the transformed covariance and cross-covariance matrices of both the UT and the SUT; 2) multiple UKF definitions; 3) issue with some reduced sets of sigma points described in the literature; 4) the conservativeness of the SUT; 5) the scaling effect of the UT on both its transformed covariance and cross-covariance matrices; and 6) possibly ill-conditioned results in SRUKF’s. With the proposed systematization, the symmetric sets of sigma points in the literature are formally justified, and we are able to provide new consistent variations for UKF’s, as such the Scaled SRUKF’s and the UKF’s composed by the minimum number of sigma points. Furthermore, our proposed SRUKF has improved computational properties when compared to state-of-the-art methods.

Index Terms—Unscented Kalman Filter (UKF), Unscented Transformation (UT).

I. INTRODUCTION

INCE the seminal work of Julier and Uhlmann in 1995 [1], the Unscented Kalman Filter (UKF) has been an object of great interest in the control community. One of the main reasons behind its popularity is the fact that the UKF has provided better results in comparison to the Extended Kalman Filter (EKF) in numerous applications, such as state estimation for battery charging [2], plasma insulins [3], and simultaneous localization and mapping (SLAM) [4], among others. However, in contrast with the more mathematically principled approaches utilized in the Gauss-Hermite Filter (GHF) of [5] or the Cubature Kalman Filter (CKF) of [6], all known UKF formulations have had their algorithms originated by ad hoc reasoning following Julier’s intuitive principle that “it is easier to approximate a probability distribution than it is to approximate an arbitrary non-linear function” [7]. Although this intuition was suitable for the initial derivation of the filter, its lack of rigor might lead to misleading interpretations and inconsistencies. In fact, some of those have already been reported in the literature. For instance:

• the UKF of [8] presents instability problems (see [9]);
• there are problems with the estimation accuracy of the UKF of [1] (see [10], [11] and Section II-C1);
• the UKF’s of [8] and [9] are valid only for scalar systems (see [12] and Sections II-D1 and II-D2).

This work is an attempt to systematize the theory behind the UKF, its first goal being to find a formal justification for the existing UKF definitions. As a result, we unify the most popular discrete-time UKF formulations into a single framework; clarify the reasons behind the aforementioned inconsistencies and propose solutions (see Sections II-B to II-F); and derive new consistent formulations of the UKF (see Section V).

This paper is organized as follows. In Section II, we briefly review non-linear filtering theory and all known discrete-time formulations of the UKF. Next, we discuss previously-known and some new inconsistencies related to these UKF’s. From Section III to V, we propose a systematization of the UKF. We introduce the concept of sigma-representations of a random variable (rv) in Section III and present new formalizations for the Unscented Transformation (UT) and the UKF in Sections IV and V, respectively. We conclude with our final remarks in Section VI. Throughout this work, the following notations and definitions are used:

• the set of all real valued rv of length $n$ is denoted by $\Phi^n$.
• $\langle x \rangle$ is its probability density function (pdf) and $E\{X\}$ or $\bar{X}$, its expected value. If $p_X(\bar{X} + x) = p_X(\bar{X} - x), \forall x \in \mathbb{R}^n$, then $X$ is symmetric. $X \sim (m, M_2, ..., M_n)^n$ stands for a rv $X \in \Phi^n$ with mean $m$ and $i$th central moment $M_i^X = M_i, i = 2, ..., k$.
• $\otimes$ stands for the Kronecker product operator.
• $\bigotimes A_i := A_1 \otimes \cdots \otimes A_n$, and $A^{\otimes n} := \bigotimes A$.
• $\sqrt{A}$ stands for a square-root matrix of the matrix $A$.
• $[A]_{p \times q}$ stands for a block matrix consisting on the matrix $A$ being repeated $p$ times on the rows and $q$ on the columns.
• \((A)_{i,j}(x)_{j_2,j_2}\) stands for a sub-matrix of the matrix \(A\) formed by the rows \(i_1\) to \(i_2\) and the columns \(j_1\) to \(j_2\). 
• \((A)_{i,j}\) stands for the \(i\)th row and \(j\)th column element. 
• \((A)_{i,j}\) and \((A)_{j,i}\), stand, respectively, for the \(j\)th column and \(i\)th row of \(A\). If no ambiguities exist, we can write 
  \[ A_{ij} := (A)_{i,j}. \]
• \([A]\), \(A \in \mathbb{R}^{n \times m}\), is such that \([A]_{i,j} := [A]_{i,j}\), where \([\cdot]\) represents the absolute value operator.
• \text{diag}(A_1, \ldots, A_q)\), where \(A_i\) is square, stands for a block diagonal matrix with each block equal to \(A_i\), \(i = 1, \ldots, q\).
• \text{sign}(A)\), where \(A \in \mathbb{R}^{n \times m}\), is such that \((\text{sign}(A))_{i,j} := 1\), if \(i, j \geq 0\), or \((\text{sign}(A))_{i,j} := -1\), if \(i, j < 0\).
• \text{tria}\{A\} represents a lower triangularization of \(A\) (e.g. QR decomposition), and \text{cdown}\{A, B\} the Cholesky factor of \(AA^T - BB^T > 0\).
• \(\text{Tab } X[p,q,n,m]\) refers to the rows \(p\) to \(q\) and the columns \(n\) to \(m\) of Table \(X\). If Table \(X\) has only one column, then \(\text{Tab } X[n,m]\) refers to its rows \(n\) to \(m\).
• A set \(\{x_i| x_i \in \mathbb{R}^n\}\) generated by a generator \([u_1, \ldots, u_n], \lambda \leq n\), is composed by all different permutations of the elements of the vector \([u_1, \ldots, u_n, 0]_{1 \times (n-\lambda)}^T\). We use the notation \(\{x_i\} = \text{gen}\{[u_1, \ldots, u_n, 0]_{1 \times (n-\lambda)}\}\) (see details in [13]).

II. NONLINEAR AND UNSCENTED KALMAN FILTERING

The Unscented Kalman Filter is a suboptimal solution for the stochastic filtering problem of a discrete-time, dynamical system described either in the additive form

\[
x_k = f(x_{k-1} \vert -), q_k, \quad y_k = h(x_k \vert -) + r_k
\]

or, more generally, in the form

\[
x_k = f(x_{k-1} \vert -, q_k), \quad y_k = h(x_k \vert -, r_k)
\]

where \(k\) is the time step; \(x_k \in \mathbb{R}^n\) is the internal state; \(y_k \in \mathbb{R}^m\) is the measured output; and \(q_k \in \mathbb{R}^q\) and \(r_k \in \mathbb{R}^r\) are the process and measurement noises, respectively. The noise terms \(q_k\) and \(r_k\) are assumed to be uncorrelated, with \(\mathbb{E}\{q_k\} = 0\) and \(\mathbb{E}\{r_k\} = 0\), and covariance matrices (CM’s) \(Q_k\) and \(R_k\), respectively.

The stochastic filtering problem consists in finding estimates of the state \(x_k\) as new measurements \(y_k\) are acquired. Based on the output history \(y_{1:k}\), the conditional mean \(E\{x_k \vert y_{1:k}\} = \int_{x_k} x_k p(x_k \vert y_{1:k}) dx_k\) is, in general, chosen to be the estimate of \(x_k\) because \(E\{x_k \vert y_{1:k}\}\) is unbiased \((E\{x_k - E\{x_k \vert y_{1:k}\}\} \vert y_{1:k}) = 0\) and it is also an optimal solution with respect to the Minimum Variance (MV) criterion [14], [15]. For linear dynamical systems, the Kalman Filter (KF) provides the optimal solution with respect to the MV criterion, as well as other criteria, when independent Gaussian noise and initial state are considered [14], [16]. However, in the case of non-linear systems, the computation of such optimal solutions tends to be computationally intractable [14], [17], [18]. Therefore, suboptimal approaches must be sought.

Suboptimal, non-linear filters can be classified under four different criteria, at least. A first classification distinguishes the filters approximating the system’s functions around the previous estimates (local methods [19])—EKF [18], second order extended Kalman filter (SOEKF) [18], [20]—, from those that do not (global methods)—Gaussian Mixture filter [21], point-mass filter [22], Sequential Monte Carlo Filters (SMCF’s) (e.g. Particle Filters, Bootstrap Filters) [23]–[26] and Markov Chain Monte Carlo based filters (MCMCF’s) (e.g. filters using Metropolis-Hastings or Gibbs sampling) [27]. A second classification (see [11], [19]) is based on whether there is the necessity of calculating derivatives of system functions—EKF and SOEKF—or not, i.e., if the filter is derivative-free—UKF [1], [7], GHF [5], Central Difference Filter (CDF) [5], Divided Difference filter (DDF) [28], and CKF [6], [29]. A third classification considers filters for which statistics of posterior random variables (rv’s) are obtained by sampling of previous pdf’s. The sampling can be stochastic, as in SMCF’s and MCMCF’s, or deterministic, as in the UKF and the DDF. Monte Carlo (MC) methods consist, essentially, on taking a very large quantity of samples of the rv of interest in a random fashion [23]–[27], while sigma point methods consist on analytically choosing a finite number of weighted samples [30]. A fourth classification takes into account if the posterior estimate of a filter is based on a Gaussian assumption (EKF, SOEKF, UKF’s, GHF, CDF, DDF, CKF) or not (Gaussian Mixture filter, point-mass filter, SMCF’s, MCMCF’s).

Among all non-linear filters, the EKF is the most widely-known and implemented in practical applications [7], [14], [18]. It is obtained as the first order truncation of the Taylor series approximation of system’s non-linear dynamics, retaining the same prediction-correction structure as the KF for linear dynamical systems. Although several filters in the literature have been proposed in order to improve upon computational aspects related to the EKF, it was just recently that UKF’s have become noticeable as a competitive and preferable alternative [7], [30]. Among the positive features related to the UKF, one can mention:

- the computational complexities of the UKF’s and of the EKF are of the same order—\(O(n^3)\)—, but UKF’s tend to attain better estimation performance;
- the UKF is derivative-free (no need to compute Jacobians), while the EKF requires the dynamics to be differentiable. Thus, unlike the EKF, UKF’s can be used in situations where Jacobians may not always exist, such as systems with discontinuities (cf. [7]).

These and other good properties related to the UKF have become well-known since its introduction (see more details in Section II-B). However, later UKF variants have been reported to be inconsistent (see Sections II-C to II-F) and it is currently difficult to assess whether these inconsistencies are present in all UKF variants. Seeking to provide clarifications, we first review all main UKF variants in the next subsection.

A. UNSCENTED KALMAN FILTERS

All UKF variants, as in the EKF, keep the structure of the Kalman filter for linear systems of one prediction (or a priori estimation) and one correction (or update) step. This can be seen, for instance, in the UKF of [31]: consider (1) and suppose that, at time step \(k\), \(\hat{x}_{k | k-1}\) and \(P_{xx | k-1}\) are
given. Choose a real \( \kappa > -n_x \) and define, for \( 1 \leq i \leq n_x \), the weights and points

\[
\begin{align*}
    w_0 & := \frac{\kappa}{n_x+n_a}, \quad w_i = w_{i+n_x} := \frac{1}{2(n_x+n_a)}, \\
    x_i & := \hat{x}_{k|k-1} + (n_x + \kappa) \hat{P}_{xx}^{k-1|k-1}, \\
    \chi_i & := x_{i+n_x} := \hat{x}_{k|k-1} + (n_x + \kappa) \hat{P}_{xx}^{k-1|k-1}.
\end{align*}
\]

For \( 0 \leq i \leq 2n_x \), define the transformed sigma points

\[
\chi_i^{k|k-1} := f\left(\chi_i^{k|k-1},k\right), \quad \gamma_i^{k|k-1} := h\left(\chi_i^{k|k-1},k\right)
\]

and their associated statistics \(((A)(\gamma)^T)\) stands for \((A)(\gamma)^T\)

\[
\begin{align*}
    \hat{x}_{k|k-1} & := \sum_{i=0}^{2n_x} w_i \chi_i^{k|k-1}, \quad \hat{y}_{k|k-1} := \sum_{i=0}^{2n_x} w_i \gamma_i^{k|k-1}, \\
    \hat{P}_{xx}^{k|k-1} & := \sum_{i=0}^{2n_x} w_i \left(\chi_i^{k|k-1} - \hat{x}_{k|k-1}\right) \left(\gamma_i^{k|k-1} - \hat{y}_{k|k-1}\right)^T + Q_k, \\
    \hat{P}_{xy}^{k|k-1} & := \sum_{i=0}^{2n_x} w_i \left(\chi_i^{k|k-1} - \hat{x}_{k|k-1}\right) \left(\gamma_i^{k|k-1} - \hat{y}_{k|k-1}\right)^T \\
    \hat{P}_{yy}^{k|k-1} & := \sum_{i=0}^{2n_x} w_i \left(\gamma_i^{k|k-1} - \hat{y}_{k|k-1}\right)^T + R_k.
\end{align*}
\]

along with the innovation’s covariance

\[
\hat{P}_{yy}^{k|k-1} := \sum_{i=0}^{2n_x} w_i \left(\gamma_i^{k|k-1} - \hat{y}_{k|k-1}\right)^T + R_k.
\]

Finally, instantiate the KF’s correction equations

\[
G_k := \hat{P}_{yy}^{k|k-1}^{-1} \hat{P}_{xy}^{k|k-1} = \frac{\hat{P}_{yy}^{k|k-1}^{-1} \hat{P}_{xy}^{k|k-1}}{\hat{P}_{yy}^{k|k-1} - \hat{P}_{xy}^{k|k-1} G_k^T}.
\]

All popular variants of the UKF for discrete-time systems have the same prediction-correction structure of the UKF of [31], but with different expressions for (3)-(7). They can be classified according to a few different criteria.

A first criterion is related to the form of the underlying dynamical system, i.e., whether the system is described in additive form (1) or in the more general form (2). Depending on the case, the a priori rv is either the previous state estimate \( \hat{x}_{k-1|k-1} \) with estimate CM \( \hat{P}_{xx}^{k-1|k-1} \) (as in the UKF of [31] above) or the previous augmented vector estimate \( \hat{x}_{k-1|k-1}^{a} \) with dimension \( n_a = n_x + n_a + n_r \), which is defined as \( \hat{x}_{k-1|k-1}^{a} := [\hat{x}_{k-1|k-1}^{T}, \hat{Q}_{k}^{T}, \hat{P}_{xx}^{k-1|k-1}^{T}] \) with covariance matrix and square-root covariance matrix

\[
\begin{align*}
    \hat{P}_{xx}^{a,k-1|k-1} & := \text{diag}\left(\hat{P}_{xx}^{k-1|k-1}, Q_k, R_k\right), \\
    \sqrt{\hat{P}_{xx}^{a,k-1|k-1}} & := \text{diag}\left(\sqrt{\hat{P}_{xx}^{k-1|k-1}}, \sqrt{Q_k}, \sqrt{R_k}\right).
\end{align*}
\]

Although it is always possible to use the general form of the UKF for either (1) or (2), the additive form of the UKF is preferable for (1) because its algorithm is computationally less expensive than the corresponding UKF in general form. Nevertheless, as pointed out by [32], the additive version yields the same result as the augmented one only if the predicted sigma points \( \chi_{i}^{k|k-1} \) are resampled from the previous mean \( \hat{x}_{k|k-1}^{a} \) and covariance matrix \( \hat{P}_{xx}^{a,k-1|k-1} \) when performing the UT in the correction step. If correlated noises are considered, the general form for the UKF with a simple change in (8)-(9) (see [2]) should be used for either (1) or (2).

**Remark 1** Filters for system descriptions besides (1) and (2) can be easily obtained. For partially-additive dynamical systems, where either the process or the measurement equations are given in additive form and the other is given in the general form, the augmented state vector of the filter is composed only by the noise of whatever function is in general form [15]. For partially nonlinear dynamical systems, where one of these equations is linear, the linear KF equations can usually be used in the part of the UKF referring to the linear equation [33].

A second major criterion to classify UKF variants is the propagation form of CM’s in the algorithm for the filter. The UKF can be in the covariance form, where the CM’s are calculated directly, or in the square-root form, where the square-root of the CM’s are computed recursively instead of propagating the CM’s themselves.

Let us first consider the UKF algorithms in the covariance form. In these algorithms, the mean and the CM of the predicted and corrected variables are approximated using a so-called Unscented Transformation of the previous estimate. Any UT variant is based on some set of sigma points which is chosen using a given mean and CM information. Roughly, an UT approximates the joint pdf of two rv’s \( X \sim (\bar{X}, P_{XX}) \) and \( Y \sim (\bar{Y}, P_{YY}) \) related by a given function \( Y = F(X) \) by introducing the sets of weighted points (the sigma points)

\[
\sigma_X = \{\chi_i, w_i^m, w_i^m\}_{i=1}^N
\]

and

\[
\sigma_Y = \{\gamma_i, w_i^m, w_i^m\}_{i=1}^N
\]

with the following sample statistics:

\[
\begin{align*}
    \mu_X & := \sum_{i=1}^N w_i^m \chi_i, \\
    \Sigma_{XX} & := \sum_{i=1}^N w_i^m (\chi_i - \mu_X)^T (\chi_i - \mu_X), \\
    \mu_Y & := \sum_{i=1}^N w_i^m \gamma_i, \\
    \Sigma_{YY} & := \sum_{i=1}^N w_i^m (\gamma_i - \mu_Y)^T (\gamma_i - \mu_Y), \\
    \Sigma_{XY} & := \sum_{i=1}^N w_i^m (\chi_i - \mu_X) (\gamma_i - \mu_Y)^T.
\end{align*}
\]

For any UT variant, the points and weights in the sigma set \( \sigma_X \) are required⁠¹ to be such that \( \mu_X = \bar{X}, \Sigma_{XX} = P_{XX} \) and, as consequence, \( \mu_Y, \Sigma_{YY} \) and \( \Sigma_{XY} \) are expected to be, respectively, equal to \( \bar{Y} \) and \( P_{YY} \) up to their second order Taylor approximations [7]. In additive UKF’s for (1), generally two UT’s are performed: one for \( F(X) = f(X,k) + q_k \) with \( X \) being the previous estimate \( \hat{x}_{k-1|k-1} \) and another for \( H(X) = h(X,k) + r_k \) with \( X \) being \( \hat{x}_{k|k-1} \). For general UKF’s, one has UT for \( F(X) \) with \( X = \hat{x}_{k|k-1} \) and for \( H(X) \) with \( X = \hat{x}_{k|k-1}^{a} \).

Variants of the UKF can be classified according to the underlying UT used. The UT variants differ from each other depending on how the transformed statistics are calculated and by their choice of the sigma set \( \sigma_X \). All basic sigma set variants in the literature are presented in Table I. The ones from [1] and [7] are equivalent if one chooses \( w_0 = \kappa/(\kappa + n) \) (cf. Tab I [1:2]). Hence, we can say that the general and additive UKF’s of [1] and [7] are equivalent (cf. Tab III [1:2:1:3] and Remark 3). The basic difference is in their choice of the tuning parameter \( w_0 \) or \( \kappa \).

⁠¹In the literature, these requirements are said to be properties of UT’s. In Section II-C, we present several counter-examples to these claims.
TABLE I

<table>
<thead>
<tr>
<th>Sets of Sigma Points for $X \sim (X, P_{XX})^n$</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Symmetric set of [1]</strong> ($N = 2n+1$). Choose $\kappa &gt; 0$. Set $w_0 = \frac{1}{2N+n}$, $X = \bar{X}$ and, for $i = 1, ..., n$: $w_i = w_{i+n} = \frac{1}{2(n+n)}$. $\Delta X_i = \left(\sqrt{(n+1)P_{XX}}\right)^T \chi_i = X + \Delta X_i$; $\chi_{i+n} = X - \Delta X_i$.</td>
</tr>
<tr>
<td><strong>Symmetric set of [2]</strong> ($N = 2n+1$). Choose $w_0 &lt; 1$. Set $X_0 = X$, and, for $i = 1, ..., n$, $w_i = w_{i+n} = \frac{1}{2n}$, $\Delta X_i = \left(\sqrt{\frac{1}{w_0}P_{XX}}\right)^T \chi_i = X + \Delta X_i$; $\chi_{i+n} = X - \Delta X_i$.</td>
</tr>
<tr>
<td><strong>Reduced set of [8]</strong> ($N = n+1$). Choose $0 \leq w_0 \leq 1$. Define $w_0 = w_1 = \frac{2}{n+1}$, $w_i = 2^{-i} w_{i-1}$ for $i = 2, ..., n+1$; $\chi_0 = [0]$ (\chi_i = \left[-1/2\sqrt{n}\right]^T \chi_i = \chi_i) for $i = 1, ..., n$; $\chi_{i+n}^+ = \left[0_i, 0_i^T \chi_i^+\right] = \left[0_i, 0_i^T \chi_i^+ + \bar{X}\right]$.</td>
</tr>
</tbody>
</table>
| **Spherical simplex set of [9]** ($N = n+2$). Choose $0 \leq w_0 \leq 1$. Set $w_0 = w_1 = \frac{n+1}{n}$, $i = 1, ..., n+1$; $\chi_0 = [0]$; $\chi_i = \chi_i^-$ \(\chi_i = \chi_i^+\) for $1 < i < n$; $\chi_{n+1} = \chi_{n+1}^+ = \chi_{n+1}^- = \chi_{n+1}$ \(\chi_{n+1} = \chi_{n+2}^+ = \chi_{n+2}^-\) \(\chi_{n+2} = \chi_{n+2}^+ = \chi_{n+2}^-\) \(\chi_{n+2} = \chi_{n+2}^+ = \chi_{n+2}^-\).

TABLE II

<table>
<thead>
<tr>
<th>Other Uncentered Transformations</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Scaled UT [38]</strong></td>
</tr>
<tr>
<td>1. (\chi_i, w_i^{\alpha} N^{-1\alpha}, \alpha \geq 0, 2 \leq i \leq N; w_i = w_i^{\alpha} = w_i^{\alpha - 2} - 2 - \alpha^{-2}; \chi_i = \chi_i^{\alpha} - \chi_i^{\alpha - 2}; w_i = w_i^{\alpha} - 2).</td>
</tr>
<tr>
<td>2. (\Gamma_i, w_i^{\alpha} N^{-1\alpha}, \alpha \geq 0, 2 \leq i \leq N; w_i = w_i^{\alpha} = w_i^{\alpha - 2} - 2 - \alpha^{-2}; \chi_i = \chi_i^{\alpha} - \chi_i^{\alpha - 2}; w_i = w_i^{\alpha} - 2).</td>
</tr>
<tr>
<td>3. (\Sigma_i, w_i^{\alpha} N^{-1\alpha}, \alpha \geq 0, 2 \leq i \leq N; w_i = w_i^{\alpha} = w_i^{\alpha - 2} - 2 - \alpha^{-2}; \chi_i = \chi_i^{\alpha} - \chi_i^{\alpha - 2}; w_i = w_i^{\alpha} - 2).</td>
</tr>
<tr>
<td><strong>Auxiliary form of the UT (AuxUT) [38]</strong></td>
</tr>
<tr>
<td>4. (g(X, x, \alpha, \kappa) \sim \kappa^{-1} \mathcal{N}(\mathcal{F}(x) + \mathcal{F}(c) + \mathcal{F}(c), \mathcal{F}(c))).</td>
</tr>
<tr>
<td>5. (\Gamma_i, w_i^{\alpha} N^{1\alpha}, \alpha \geq 0, 2 \leq i \leq N; w_i = w_i^{\alpha} = w_i^{\alpha - 2} - 2 - \alpha^{-2}; \chi_i = \chi_i^{\alpha} - \chi_i^{\alpha - 2}; w_i = w_i^{\alpha} - 2).</td>
</tr>
<tr>
<td>6. (\Sigma_i, w_i^{\alpha} N^{1\alpha}, \alpha \geq 0, 2 \leq i \leq N; w_i = w_i^{\alpha} = w_i^{\alpha - 2} - 2 - \alpha^{-2}; \chi_i = \chi_i^{\alpha} - \chi_i^{\alpha - 2}; w_i = w_i^{\alpha} - 2).</td>
</tr>
</tbody>
</table>

One can consider the geometrical distribution of the sigma points. For points $\chi_i$ symmetrically distributed in a geometrical sense, we have the UKF’s of [1], [7], [35], and [36]. For asymmetric sets, one can mention the reduced of [8]; the spherical simplex of [9]; the simplex of [34]; and the minimum of [12].

Another classification criterion considers the number of sigma points in the sigma set $\sigma_X$ for the previous variable $X \in \Phi^m$. As can be seen by Table I, among the major general and additive UKF variants that use symmetric sets, the UKF’s of [1], [7], and [35] use $\sigma_X$ with $N = 2n+1$ points. Among the major general and additive UKF variants that use asymmetric sets, the UKF’s of [9] use $\sigma_X$ with $N = n + 2$ points and the UKF’s of [8], [34], and [12] use $\sigma_X$ with $N = n + 1$ sigma points, which is the minimum amount [8]. Finally, the UKF of [36] uses $\sigma_X$ with $N = 2n^2 + 1$ points and the one of [37] uses $N = \phi^m, \phi \in \mathbb{N}$. One can consider which moments of the sigma set $\sigma_X$ are equal to the moments of the previous rv $X$. The sets $\sigma_X$ of the UKF’s of [1], [7], [8], [12], [34], [35], and [36] match the first two moments of any $X$. If $X$ is symmetric, then the sets $\sigma_X$ of the UKF’s of [1], [7], [35] and [36] match also all odd moments. If $X$ is Gaussian, then $\sigma_X$ in the UKF of [36] further matches the fourth moment of $X$ and the one of [37] matches the first two moments of $X$. However, not all UKF variants have their $\sigma_X$ matching the first and second moments of $X$ (see Section II-D).

Considering how the transformed statistics are calculated, one can distinguish two more UT’s other than the first six that arrive from the different choices of $\sigma_X$ (the ones of Table I): the AUUT of [38] and the scaled UT of [38] (cf. Table II). Their sigma sets are scaling transformations of a previous sigma set and the statistics calculations are modified versions of (11)-(13) (cf. Table II). According to [38], the auxiliary and the scaled forms yield the same results. Note that [39] presented an embryonic form of these scaling unscented forms.

In summary, all additive UKF’s in covariance form in the literature are represented in Table III using the sigma sets of Table I and the UT’s of Table II and (10)-(13) [all UKF variants for the general case (2) can be obtained similarly with a corresponding slightly modified version of Table III].

**Remark 2** Some particularities of Tables I, II, and III should be highlighted. If, for example, one wished to implement the UKF of [31], it would be necessary to choose $w_i^{m} = w_i$ when calculating any mean and $w_i^{a} = w_i$ when calculating any covariance matrix. As an example of how these tables can be used, the additive form UKF of [31], (3)-(6), can be obtained by taking Tab III [1,1,3]. The previous set of this filter is the symmetric set of [1] calculated for $X = \bar{x}_{k-1|k-1}$ and $P_{XX} = E_{xx}^{k-1|k-1}$.

**Remark 3** Due to the difficulty of describing UKF’s as presented in the original formulations in a simple and systematized way, the forms of the UKF’s shown in Table III are not necessarily the ones introduced by their corresponding authors. Nevertheless, the forms contained in this table, different from the original ones, are usually extensions (e.g., the additive form for the symmetric UKF of [7] in Table I is slightly more general). Besides, some of these extensions have already been explicitly proposed (e.g., the additive form of the symmetric UKF of [11] was modified in [31]).
These variants are not equivalent, one cannot properly point out analyze some of their properties. The reduced set of [8], Tab I [3].

SRUKF of [4]; SRUKF of [4] (general form (2) with the set of [1], Tab I [7], Tab II [2], Tab II [3]) – it has a scaling parameter \( \lambda \). In order to propagate square-root matrices in square-root UKF’s, one should use algorithms that are less susceptible to such errors, which is the case of KF’s in square-root form [42]. In order to present these UKF’s because, instead of presenting their expressions, these works only show procedures from which these UKF’s can be obtained.

Remark 4 There are three other UKF’s that are not presented in Table III: [40] describes a symmetric UKF matching up to the 4th central moment of the previous rv; [39], an asymmetric UKF matching up to the 3rd central moment of the previous rv; and [41], a symmetric UKF matching up to the 8th central moment of a scalar Gaussian rv. Table III does not show these UKF’s because, instead of presenting their expressions, these works only show procedures from which these UKF’s can be obtained.

Remark 5 Although some sigma sets can be composed by negative weights, one should note that these can lead to non-positive sample CM’s [7] and to higher errors [13].

There are applications where the machine precision is such that rounding errors can cause KF’s to diverge. Hence, one should use algorithms that are less susceptible to such errors, which is the case of KF’s in square-root form [42]. In order to propagate square-root matrices in square-root UKF’s, one can use the QR decomposition and Cholesky update factor techniques (see [43]). To date, we are aware of five variants for the Square-Root Unscented Kalman Filters (SRUKF’s): SRUKF of [43] (system in additive form 1) with the sigma set of [35], Tab I [7], and statistics calculation (10)-(13)); SRUKF of [30] (general form 2) with the set of [35]; SRUKF of [44] (additive form 1) with the spherical simplex set of [9], Tab I [4]); SRUKF of [4] (general form 2) with the set of [11], Tab I [11]); the Improved SRUKF of [45] (additive form 1) with the reduced set of [8], Tab I [3]).

Next, we present general comments about UKF variants and analyze some of their properties.

B. Definitions for UKF’s

1) Variations on UKF definitions: From Section II-A, it is clear that there are many UKF variants. Given that, in general, these variants are not equivalent, one cannot properly point out which one is “the definition” for the UKF. Nonetheless, most works in the literature use the term ‘UKF’ when referring to either the UKF of [1] as can be seen in [6], [19] – or to the UKF of [35] as can be seen in [11], [46]. By comparing their sets of sigma points (cf. Tab I [1] with Tab I [7]), one can see that there are two main differences between these filters. First, the UKF of [1] uses \( \alpha \) to calculate the weights and the sigma points, while the one of [35] uses a term \( \lambda = \alpha^2 (n + \kappa) - n \) to do so. Second, in the UKF of [35], \( \omega_0^m \) and \( \omega_2 \) are distinct objects, while in the UKF of [1], \( \omega_0^m = \omega_2 \).

2) Variation on scaled UKF definitions: Although the UKF of [35] (Tab III [5, 1-3]) is described and widely referred to as a non-scaled UKF (cf. [11], [46]), Merwe himself, one of the authors in [35], describes this filter as a scaled UKF form (cf. [30]) – it has a scaling parameter \( \alpha \) (cf. Tab I [7]). Apart from that, one should notice that this scaled UKF form differs from the ones proposed by [38] (the ones using the U.T’s of Table II).

C. Accuracy of the UKF’s

1) Transformed covariance matrix: As [11] states, a large number of papers repeat the statement of [7] that the CM of transformed rv’s is matched up to the second order in the UKF’s when the mean and the CM of the prior rv are matched [7]. However, that is not true for all UKF variants. Indeed, [11] has already pointed out this issue for the symmetric UKF of [35] by providing a counter-example: considering \( X \sim N(0, n_1, I_n) \) and \( Y := f(X) = X^T X \), the theoretical result for the CM of \( Y \) is \( P_{yy} = 2n \). The second order Taylor approximation gives the correct result, but the UT of [35] differs from it (see Table II in [11]).

2) Transformed cross covariance matrix: The transformed cross-covariance matrix (CCM) is necessary for the UKF, but the literature has not yet provided an estimation quality for it.

D. Sigma sets composed by less than 2n sigma points

1) Reduced set of [8]: This set has two drawbacks. First, it can be numerically unstable for great values of \( n \) due to the fact that the weights are composed by fractions of \( 2^n \) [9]. Second, neither the sample mean, \( \mu_n \), nor the sample CM, \( \Sigma_{XX} \), are equal to the mean and CM of the prior distribution when \( n \) is greater than one [12]. In fact, from Tab I [3], for \( n = 2 \), \( X \sim ([0]_{2 \times 1}, I_2)^2 \), \( w_0 = 0.5 \), and using (11) with \( w_i^e = w_i^m = w_i \), we have that \( \mu_X = 1/4[-1, 3]^T \neq [0]_{2 \times 1} \) and \( \Sigma_{XX} = \frac{1}{16} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 1 & 5 & 1 & 5 \\ 1 & 5 & 1 & 5 \\ 1 & 5 & 1 & 5 \end{bmatrix} \neq I_2 \).

2) Spherical simplex set of [9]: This set does not present the instability problem of the set of [8], but still has the same problem that neither \( \mu_X \), nor \( \Sigma_{XX} \), are equal to the mean and the CM of \( X \) when \( n \) is greater than one [12]. In fact, from Tab I [4], for \( n = 2 \), \( X \sim ([0]_{2 \times 1}, I_2)^2 \), \( w_0 = 0.5 \), and using (11) with \( w_i^e = w_i^m = w_i \), we have that \( \mu_X = 1/(2\sqrt{6})[0, 1]^T \neq [0]_{2 \times 1} \) and \( \Sigma_{XX} = \text{diag}(1, 53/96) \neq I_2 \).

3) Minimum set of [34]: The CM for this set is not matched. Using Tab I [5] and (11), it can be shown that \( \Sigma_{XX} = P_{XX} + 1/(n + 1) \begin{bmatrix} X_1 \end{bmatrix}_{1 \times (n+1)} [X_1^T]_{(n+1) \times 1} \) for \( X \sim (X, P_{XX})^n \). In fact, for \( X \sim (3, 4)^4 \), \( \Sigma_{XX} = 13 \neq 4 \).
4) Minimum set of [12]: This is the only sigma set composed by less than 2n points matching the mean and CM of X.

E. Scaled Transformations

1) Scalable sigma sets: the authors in [38] state that the “scaled uncented transform [...] allows any set of sigma points to be scaled by an arbitrary scaling factor” (the italic is of the authors and the bold is ours). However, suppose X = ([0,2]×1, I3) and that the previous sigma set is \( \chi_1 = \left[ \sqrt{2}, 0 \right]^T \), \( \chi_2 = \left[ 0, \sqrt{2} \right]^T \), \( \chi_3 = \left[ -\sqrt{2}, 0 \right]^T \), and \( w_1 = w_2 = w_3 = w_4 = \frac{1}{4} \). From Tab II [1:3], one can see that, for \( \alpha = 0.5 \) and \( \chi_1 = [0, 2]^T \), the sample mean, \( \mu_X \), and the sample CM, \( \Sigma_X \), of the (previous) scaled sigma set are \( \mu_X' = [-\sqrt{2}, 0]^T \neq [0, 2]^T \) and \( \Sigma_X' = \text{diag}([-3, 1]^T) \neq I_2 \).

This example shows that the sample mean and the sample CM of \( \chi' \) are not equal to the mean and CM of X, respectively. In fact, as one can see from the following theorem, this property is not guaranteed to hold for any sigma set, except for those having one sigma point equal to the mean of X.

Theorem 1 Consider X ∼ (X, P_{XX})\( ^n \) and a function f : \( \mathbb{R}^n \rightarrow \mathbb{R}^n \) defining a new rv Y := f(X) and consider a set of sigma points \( \chi = \{ x_i, w_i \}_{i=1}^N \) for X. Consider also the set of scaled sigma points \( \chi' = \{ \chi_i', w_i' \}_{i=1}^N \) obtained from the scaled UT of [38] (Tab II [1:3]). We have that:

1) if \( \sum_{i=1}^N w_i = 1 \), then \( \sum_{i=1}^N w_i' = 1 \);
2) \( \mu_{X'} = \frac{1}{\alpha} \mu_X + \chi_1 (1 - \frac{1}{\alpha}) \);
3) if \( \chi_1 = \bar{X} \), then \( \mu_{X'} = \mu_X \);
4) if \( \alpha \neq 1 \), then \( \chi_1 = \bar{X} \Leftrightarrow \mu_{X'} \neq \mu_X \);
5) if \( \chi_1 = \bar{X} \), then \( \Sigma_{X'} = \Sigma_X \).

Proof: Suppose \( \sum_{i=1}^N w_i = 1 \), then \( \sum_{i=1}^N w_i' = \frac{1}{\alpha} (1 + 1 - \frac{1}{\alpha}) = 1 \). For the second and third assertion, \( \mu_{X'} = \frac{1}{\alpha} \mu_X + \chi_1 (1 - \frac{1}{\alpha}) \), which, supposing \( \chi_1 = \bar{X} \), gives \( \mu_{X'} = \mu_X \) and, supposing \( \mu_{X'} = \mu_X \), \( \alpha \neq 1 \), gives \( \chi_1 = \bar{X} \).

The last assertions can be proven by the fact that \( \sum_{X'} = w_1' (\chi_1 - \bar{X}) (\bar{X})^T + \sum_{i=2}^N w_i' (\chi_1 + \alpha (\chi_1 - \bar{X}) - \mu_X) (\bar{X})^T \), which, for \( \chi_1 = \bar{X} \), gives \( \Sigma_{X'} = \Sigma_X \).

Therefore, the scaled UT of [38] is restrictive in the sense that this UT does not provide the mentioned results for any previous set of sigma points. On the other hand, this restriction is not applied to the AuxUT of [38].

2) Covariance matrix: Consider X ∼ N([0,1]×1, I3)\(^3\) and Y := f(X) = X\(^T\)X. Then, \( \bar{Y} = 3 \) and \( P_{YY} = 6 \). Using the scaled UT of [38] with the symmetric sigma set of [7], we get, from (12), \( \mu_Y = 3 = \bar{Y} \) and \( \Sigma_{YY} = 3\alpha^2 - 3 \neq P_{YY} \). This result shows two problems involving the matching of the CM. First, the transformed CM for this scaled UT is not matched up to the order 2, but only to the order 1. Second, the scaling factor modifies the CM even for second order polynomial approximation.

3) Cross-covariance matrix: Similar to the case for the non-scaled UTs, the estimation quality of CCM’s for the scaled UT of [38] and for the AuxUT of [38] has not been presented in the literature yet. Besides, there is no mention of the influence of the scaling factor on the transformed CCM for the UKF of [35]. Since it is desirable to match the first and the second moments, the free parameter \( \alpha \) should modify only the third and higher terms. However, consider X ∼ N([0,1]×1, I3) and Y := f(X) = X\(^T\)X. Then, from Tab I [7], (12) and (13), we have \( \mu_Y = 3 = \bar{Y} \) and \( \Sigma_{YY} = 6\alpha [I_3]_{\alpha} - \frac{9}{2\alpha} [I_3]_{\alpha} \). Therefore, the second order term is also modified.

F. Square-root forms of the UKF’s

1) Downdating the Cholesky factor: For an equation in the form AA\(^T\) = RR\(^T\) - SS\(^T\), where A, R, S are Cholesky factors, we say that A is a downdated Cholesky factor of R by S. There are three parts within the SRUKF algorithms in the literature where Cholesky factors are downdated: in the calculations of the square-root matrices (SRM’s) of the predicted state’s CM, of the innovation’s CM, and of the corrected state’s CM. In the first two steps, the downdating steps are performed only for the sigma points with negative weights, while, in the last, they are always performed.

Since the direct downdating of a Cholesky factor is “inherently more ill-conditioned than if Q (the Q matrix of a QR decomposition) is also available” [47] (the comment within parentheses is ours), filters resulting from the substitution of downdating steps by QR decompositions – or, more generally, by any triangulation technique [6] – should be computationally more stable. In fact, [48] has developed such a technique for calculating the SRM of the corrected state’s CM for quadrature Kalman filters and [6] for the CKF.

2) Square-Root Scaled UKF: The literature does not present any filter conjugating the SRUKF with the scaled UT of [38] (Tab II [1:3]) nor with the AuxUT (Tab II [4:6]).

3) Square-Root UT: Although there are definitions for SRUKF’s, we have not been able to find any definition for a Square-Root Uncented Transformation (SRUT). The importance of defining a SRUT can be motivated by the importance of defining an UT: it gives its resulting UKF a better mathematical formal principle; it is possible to study a UKF focusing on its respective UT, since it is the core difference between the UKF relative to other nonlinear KF-based filters; one can apply a UT not only within the KF framework, but in any framework or application that requires uncertainty propagation (e.g. [49]) or within other stochastic filter (e.g. [50]).

G. Conclusion of the literature review

This section showed that many variations for the UKF have been proposed. Some of them are different in the recursive filter framework, others in the composition of sigma sets, and others in the transformation from the previous to the posterior sigma sets. Sections II-B to II-F presented some problems and gaps within the UKF theory.

The difficulty in gathering all results related to Unscented and the problems within some of them reveal a lack of foundation in terms of mathematical principles and also the absence of mathematical solutions generalizing the sigma sets, UTs and UKFs of the literature. In order to address these needs, we propose a systematization that treats the construction of UKF’s in parts. We first consider the problem of estimating the mean of a non-linear transformation, which will lead us to the definition of the sigma-representation.
III. SIGMA-REPRESENTATIONS

Given a rv $X \in \Phi^n$ with pdf $p_X(x)$, many problems, such as calculating the moments of an rv, can be reduced to the problem of finding the posterior expectation

$$E\{f(X)\} = \int_{\mathbb{R}^n} f(\xi)p_X(\xi)d\xi,$$

(14)

for an appropriate function $f : \mathbb{R}^n \rightarrow \mathbb{R}^{n\prime}$. As a first attempt to solve this problem, one could consider using numerical integration techniques. In the scalar case ($n = n_y = 1$) and if the function $f$ is well approximated by a polynomial of order $2N - 1$, the Gaussian quadrature methods give approximate solutions for (14) of the form (see [5], [51]–[54])

$$E\{f(X)\} = \int_{-\infty}^{\infty} f(\xi)p_X(\xi)d\xi \approx \sum_{i=1}^{N} w_i f(x_i),$$

where $x_1, \ldots, x_N$ are samples of $X$. For $X$ being a standard scalar normal rv, the solution is obtained by the Gauss-Hermite Quadrature (GHQ) [5], [30], [51]–[53], [55]. The multivariate case can be obtained by first using a stochastic decoupling technique $X' = \sqrt{P_{XX}}^{-1}(X - \bar{X})$, where $X'$ is a multivariate standard Gaussian rv. Then, for $\tilde{f}(X) := f(\sqrt{P_{XX}}^T X + \bar{X})$, the GHQ is applied on the form

$$E_{X'}\{\tilde{f}(X')\} = \int_{\mathbb{R}^n} \tilde{f}(\xi)p_X(\xi)d\xi \approx \sum_{i=1}^{N} w_{i_1} \cdots w_{i_n} \tilde{f}(x_{i_1}, \ldots, x_{i_n}),$$

and $E_X\{f(X)\}$ is obtained from $E_{X'}\{\tilde{f}(X')\}$ [5]. An alternative to solving the multivariate Gaussian case is to use the spherical cubature rule with the Gaussian Quadrature after performing a Cartesian-to-spherical coordinate transformation. In fact, consider the Gaussian case $p_\rho(\xi) = \exp\{-\xi^T \rho \xi\}$ and let $\xi = \rho y$, with $y^T y = 1$, $\rho \in [0, \infty)$. In this case, (14) becomes

$$E\{f(X)\} = \int_{U_\rho} S(\rho)\rho^{n-1}\exp(-\rho^2)d\rho, \quad S(\rho) := \int_{U_n} f(\rho y)d\phi(y),$$

where $U_n := \{u \in \mathbb{R}^n | u^T u = 1\}$ and $\phi(\bullet)$ is the spherical surface measure of $U_n$ [6]. The spherical integral $S(\rho)$ is solved by the spherical cubature rule, while the expectation by a Gaussian Quadrature rule [6].

Instead of using a quadrature solution, one can obtain a suboptimal solution by approximating the function $f$. For instance, one can use linearization or higher-order polynomial approximations of the kind $f(x) \approx \sum_k a_k x^k$ [56]. In this case, (14) would be approximated by $E\{f(X)\} = \sum_k a_k \int_{\mathbb{R}^n} \xi^k p_X(\xi)d\xi$. Well-known methods are the trapezoidal rule, Simpson’s Rule, the Newton-Cotes Formulas, the Clenshaw-Curtis Integration, among others [56].

Another alternative for obtaining (14) is by approximating $p_X(x)$. We can classify this type of suboptimal approximation into two categories, namely Monte Carlo methods [23]–[27], [57] and sigma-point methods [30], [55]. Monte Carlo (MC) methods consist of taking a very large quantity of samples $x_i$ of $X$ (the method gets more accurate as the number of samples $N \rightarrow +\infty$) randomly [23], [24], [26], [27]. Sigma point methods, on the other hand, consist on analytically choosing finite $N$ samples $x_i$ and weights $w_i$ [30]. These approaches can be viewed as generalized—negative weights are admitted—discrete approximations of $p_X(x)$.

There is some overlap in this type of classification, as well as other interpretations. Some sigma-point formulas can be obtained from integration approaches [15], [33], [58]. For instance, [6] derives a particular case of the symmetric sigma-point set of [7] (Tab I [2]) using the spherical cubature quadrature; and [26] and [13] derive the fifth-order sigma-point set (Tab I [8]) also by this quadrature rule [15]. It is worthwhile to mention that the symmetric sigma-point set of [7] can also be viewed as a statistical linear regression technique [31].

All techniques for expected value calculation can be used for dynamical systems (1) or (2) in order to obtain recursive filters. For instance, GHQ yields the GHF [5] when applied in a KF framework; the cubature spherical rule yields the CKF [6], [29]; the Central Difference technique, the CDF [5]; the linearization and the second order approximation of the functions yield the EKF and the SOEKF, respectively; different UT's yield different forms of the UKF; Stirling’s interpolation formula yields the Divided Difference Filter (DDF) [28]; and the Monte Carlo methods yield SMCF's (e.g. PF's [23]–[26]) or MCMCF's [27].

The DDF and the CDF are considered to be “essentially identical” [30]. The CKF is a particular case of the derivations in [36] and [13], where the CKF is also showed to be equivalent to the UKF of [1] (Tab III [1,1:3]) by making the central weight equal to zero [11], [19].

The UKF of [1] is showed to be a particular case of the GHF in the scalar case ($n = n_y = 1$) [5]. In fact, consider a scalar standard normal rv $X \sim N(0,1)$. Both a GHQ approach of order $N = 3$ and a sigma set of [1] with $k = 2$ and $n = 1$ would yield the set $\chi_1 = \sqrt{3}, \chi_2 = 0, \chi_3 = \sqrt{3}, w_1 = w_3 = \frac{1}{2}, w_2 = \frac{1}{3}$ [5]. However, for larger lengths of the state vector, this equivalence does not hold. The GHF is $O(n^3)$, while the UKF of [1] is $O(n^5)$ [5], [13]. In fact, for $X = \{(0)_{2\times1}, I^T_{2\times2}\}$, the Gauss-Hermite set would be $w_i = 1/3, i = 1, 3, 7, 9; w_j = 1/9, j = 2, 4, 6, 8; w_5 = 4/9; \chi_1 = \chi_3 = -\chi_7 = -\chi_9 = \sqrt{3}, \sqrt{3}^T; \chi_2 = -\chi_8 = [0, \sqrt{3}]^T; \chi_4 = -\chi_6 = [\sqrt{3}, 0]^T; \chi_5 = [0]_{2\times1}$; while the sigma set of [1] (Tab I [1]) for $k = 2$ and $n = 2$ would be $w_0 = 0.5; w_1 = \cdots = w_4 = 0.125$; and $x_0 = [0, 0]^T; x_1 = -x_3 = [2, 0]^T; x_2 = -x_4 = [0, 2]^T$.

In order to properly construct the systematization of the UKF filtering theory, we propose definitions of three fundamental mathematical elements: (I) the sigma ($\sigma$)-representation; (II) the Unscented Transformation; and (III) the recursive filters. The first is an approximation of a rv’s pdf by a set of weighted points via moment matching. We say
that a set is an lth order σR of a rv if the central moments of its samples are equal to the central moments of the chosen rv up to, and including, order l.

The case l = 2 is of particular interest, since the majority of works in Unscented literature focus on second order moment matching [11]-[4], [7]-[9], [30], [38], [40], [43], [59]. This is mainly motivated by three facts. First, these are usually the estimated statistics within a stochastic filter. Second, they fully describe a Gaussian distribution [55]. Third, the mean is the point estimate with the least mean squared error.

The notation $M^j_X$ stands for the jth central moment of $X \in \Phi^n$, and is defined, for even and odd j, respectively, as

$$M^j_X := E \left\{ \left[ (X - \bar{X}) (X - \bar{X})^T \right]^{\otimes \frac{j}{2}} \right\},$$

$$M^j_X := E \left\{ \left[ (X - \bar{X}) (X - \bar{X})^T \right]^{\otimes \frac{j-1}{2}} \otimes (X - \bar{X}) \right\}.$$

**Definition 1 (Sigma (σ)-Representation)** Define, for even and odd j, respectively,

$$M^j_X := \sum_{i=1}^N w^{(j)}_i \left( \chi_i - \mu_X \right) \left( \chi_i - \mu_X \right)^T \otimes \frac{j}{2},$$

$$M^j_X := \sum_{i=1}^N w^{(j)}_i \left( \chi_i - \mu_X \right) \left( \chi_i - \mu_X \right)^T \otimes \frac{j-1}{2} \otimes \left( \chi_i - \mu_X \right),$$

as the jth sample central moment of $\chi := \{X_i, w_1^{(1)}, \ldots, w_i^{(l)} \}_{i=1}^N$, which has sample mean $\mu_X := \sum_{i=1}^N w^{(1)}_i \chi_i$ and consider $X \sim (\bar{X}, M^2_X, \ldots, M^N_X)^n$. Then $\chi$ is an lth order N points σ-representation (lthNσR) of X if

$$w^{(j)}_i \neq 0, i = 1, \ldots, N \text{ and } j = 1, \ldots, l; \mu_X = X; M^j_X = M^j_X, j = 2, \ldots, l.$$

Besides, assume $\chi$ to be an lthNσR of X, then:

- $\chi$ is normalized if $\sum_{i=1}^N w^{(j)}_i = 1$, $j = 1, \ldots, l$.
- $\chi$ is symmetric if, for N odd, $\chi_i - \chi_N = - (\chi_i + \chi_N)$ and $w^{(j)}_i = w^{(j)}_{N+1-i}$, $1 \leq i \leq \frac{N-1}{2}$; or for N even, $\chi_i - \chi_N = - (\chi_i + \chi_N)$ and $w^{(j)}_i = w^{(j)}_{N-i}$, $1 \leq i \leq \frac{N}{2}$.
- $\chi$ is homogeneous if, for N odd, $w^{(j)}_i = w^{(j)}_{N+1-i}$, $1 \leq i \leq N-1$, or for N even, $w^{(j)}_i = w^{(j)}_{N-i}$, $1 \leq i \leq N$.

When calling an lthNσR of X, the reference to the lth order can be omitted if $l = 2$. Also, the reference to N point and/or to X can be omitted in case they are obvious from the context or irrelevant for a given statement.

**Theorem 2** $X \sim (\bar{X}, M^2_X, \ldots, M^N_X)^n$ admits a normalized lthNσR if and only if there exists $E \in \mathbb{R}^{n \times N}$, $w^{(1)} := \begin{bmatrix} w^{(1)}_1 & \cdots & w^{(1)}_N \end{bmatrix}^T$, $W^{(j)} := \text{diag}(w^{(j)})$, where $w^{(j)} := \begin{bmatrix} w^{(j)}_1 & \cdots & w^{(j)}_N \end{bmatrix}^T$, for $j = 2, 3, \ldots, l$ that satisfy, for l even, the following equations:

$$E w^{(1)} = 0;$$

$$[1]_{1 \times N} w^{(j)} = 1, \quad j = 1, 2, \ldots, l.$$  (17)

If (15) - (18) admits a solution $(E, w^{(1)}, W^{(2)}, \ldots, W^{(l)})$, then a normalized lth order lthNσR of X is $\{X_i, w^{(1)}_i, \ldots, w^{(l)}_i \}_{i=1}^N$ such that $\{X_1, \ldots, X_N\} := E + [\bar{X}]_{1 \times N}$.

**Proof:** Define $E := [\chi_1 - \mu_X, \ldots, \chi_N - \mu_X]$. So, from the definition of $M^j_X$, for $j$ even,

$$M^j_X = \sum_{i=1}^N \left[ E_n^{\otimes 2} \right] w^{(j)}_i \left[ E_n^{\otimes 2} \right]^T = \left[ E_n^{\otimes 2} \right] w^{(j)} \left[ E_n^{\otimes 2} \right]^T,$$

which proves (15). (16) can be proven similarly and the remaining is trivial.

Results for odd l can be obtained in a similar fashion. The next corollary uses Theorem 2 to obtain two novel results: the minimum amount of σ-points for both the symmetric and the non-symmetric case when $P_{XX} \geq 0$. Note that the previously-stated result that the minimum number is $n + 1$ for $P_{XX} > 0$ [7]-[9] is a particular case of Corollary 1.

**Corollary 1** Let $\chi := \{X_i, w^{(1)}_i, \ldots, w^{(l)}_i \}_{i=1}^N$ be an lthNσR of $X \in \Phi^n$, which has CM $P_{XX}$ with rank $r \leq n$. Then

1) $N \geq r + 1$. If $N = r + 1$, then $\chi$ is a minimum lthNσR of X.

2) If $\chi$ is symmetric, then $N \geq 2r$. In this case, if $N = 2r$, then $\chi$ is a minimum symmetric lthNσR of X.

**Proof:** To prove assertion 1), consider, first, $E \in \mathbb{R}^{n \times N}$ and the singular value decomposition of $P_{XX} = USV^T$, where $S := \text{diag}([\alpha_1, \ldots, \alpha_r][0_{(n-r) \times 1}]^T)$, and $\alpha_1, \ldots, \alpha_r$ are the singular values of $P_{XX}$. Assume, for contradiction, $m := \text{rank}(E) < r$. Then there exists $v := \begin{bmatrix} v^T_1 & \cdots & v^T_{n-r} \end{bmatrix}^T$, $v_1 \in \mathbb{R}^r$, $v_1 \neq 0$, such that $v^T E = 0$. Then, from (15),

$$E^{(2)T} = V_1 U_2 U_3^T \iff v^T_1 V_1 S V_1 = 0 \iff v^T_1 = 0,$$

which is a contradiction. Therefore, $m \geq r$. Second, suppose $N = m$. Then, $E$ is full column rank and, from (17), $w^{(1)} = 0$, which is a contradiction for, from Definition 1, $w^{(1)} \neq 0$. So $N \geq \text{rank}(E) + 1 \geq r + 1$. To prove assertion 2), let $\chi$ be symmetric. Then $E = [E_2, -E_2]$, where $E_2 \in \mathbb{R}^{n \times N}$. So $r \leq \text{rank}(E) = \text{rank}([E_2, -E_2]) = \min \{n, \frac{N}{2}\} \iff N \geq 2r.$

**Corollary 2** Let $\chi := \{X_i, w^{(1)}_i, \ldots, w^{(l)}_i \}_{i=1}^N$ be a normalized lthNσR of $X \sim (\bar{X}, M^2_X, \ldots, M^N_X)^n$ and consider the rv $Z = AX + b$, $A \in \mathbb{R}^{n \times N}, b \in \mathbb{R}^n$. Then, the set $\Xi := \{X_i, w^{(1)}_i, \ldots, w^{(l)}_i \} \Xi_i = AX_i + b \}_{i=1}^N$ is a normalized lthNσR of Z. In particular, we have $\mu_\Xi = AX + b$ and $\Sigma_\Xi = AP_{XX} A^T$.

**Proof:** The sample mean of $\Xi$ is $\mu_\Xi = \sum_{i=1}^N w^{(i)} (AX_i + b) = Z$. The j-th sample central moment of $\Xi$, $j = 2, 4, \ldots, l$ (l even) is

$$M^j_X = \sum_{i=1}^N w^{(i)} \left( (\chi_i - \mu_X) \otimes (\chi_i - \mu_X)^T \right)^{\otimes \frac{j}{2}} (A^T)^{\otimes \frac{j}{2}}$$
The odd case can be similarly proven.

The result used by [7], [40], [59] and others that a sigma set \( \{ x_i, w_i^m, w_i^c \}_{i=1}^{2n} \) approximating an rv \( X \sim (X, P_{XX})^n \) can be obtained by the transformation \( \chi_i = \sqrt{P_{XX}} x_i + \tilde{X} \), where \( \{ x_i, w_i^m \}_{i=1}^{2n} \) is a sigma set of an rv with mean \( [0]_{n \times 1} \) and covariance matrix \( I_n \), is a particular case of Corollary 2.

**B. Particular sigma-representations**

Motivated by Theorem 2 and Corollary 1, we now seek a \( \sigma R \) for the minimum symmetric and the minimum cases.

1) Minimum Symmetric \( \sigma \)-Representation: Let \( \chi = \{ x_i, w_i^m, w_i^c \}_{i=1}^{2n} \) be a \( \sigma R \) of \( X \sim (X, P_{XX})^n \), \( P_{XX} \geq 0 \). Considering the equations of Theorem 2, suppose \( \chi \) is minimum symmetric. Then, we have \( E = \{ E, -E \} \), where \( E \in \mathbb{R}^{n \times n} \). Define \( W := diag (\{ w_i^m, w_i^c \}^T) > 0 \). Then, from (15),

\[
\left( \sqrt{\frac{E}{2W}} \right)^{-1} P_{XX} \left( \sqrt{\frac{E}{2W}} \right)^{-1}, \frac{E}{2W}^{-1} P_{XX} + [X]_{1 \times 2n}.
\]

In addition, if \( \sum_{i=1}^{2n} w_i^c = \sum_{i=1}^{2n} w_i^m = 1 \), then \( \chi \) is normalized. Moreover, if \( W = (1/(2n))I_n \) and \( w_i := w_i^m = w_i^c \), \( i = 1, \ldots, 2n \), then \( \chi = \{ x_i, w_i \}_{i=1}^{2n} \) is a (normalized) homogeneous minimum symmetric sigma-representation of \( X \).

If an extra point located on \( X \) is added to this \( \sigma R \), neither the sample mean, nor the sample CM, will be modified. Therefore, the extra point’s weight can act as a tuning parameter.

2) Minimum Sigma-Representation: In this section, we look for a \( \sigma R \) with the least amount of (non-symmetric) sigma points. For \( P_{XX} > 0 \), this number is \( n + 1 \) (cf. Corollary 1). We first present a heuristic for finding this \( \sigma R \), followed by a formal and more general minimum \( \sigma R \) (MiSiR) in Theorem 3. At the end of this section, Corollary 5 shows that this minimum \( \sigma R \) has the minimum sigma set of [12] as a particular case.

Let \( \{ x_i, w_i \}_{i=1}^{n+1} \) be a normalized minimum \( \sigma R \) of \( X \sim (X, P_{XX})^n \). From Theorem 2 and Corollary 1,

\( E = \{ E, e \} \), where \( E \in \mathbb{R}^{n \times n} \) and \( e \in \mathbb{R}^{n \times n} \). Define \( w = [w_1, \ldots, w_n]^T \) and \( W := diag (w) > 0 \). Then, from (15),

\[ e = -w_{n+1} E \tilde{W} \]. Substituting it on (17),

\[ P_{XX} = \tilde{E} W E^T + w_{n+1} \tilde{E} w w^T E^T = \tilde{E} \sqrt{W} \left( I_n + v v^T \right) \left( \tilde{E} \sqrt{W} \right)^T, \]

where \( v := w_{n+1}^{-1/2} \sqrt{\tilde{W}} \). Then \( \tilde{E} := \sqrt{P_{XX}} \left( I_n + v v^T \right)^{-\frac{1}{2}} \left( \sqrt{\tilde{W}} \right)^{-\frac{1}{2}} \) is a sufficient condition. Moreover, from (17),

\[ w_{n+1} = 1/(1 + \sum_{i=1}^{n} v_i^2). \]

The following theorem gives the minimum \( \sigma R \) for an even more general solution (without the \( w_i > 0 \) restriction).

**Theorem 3** Minimum \( \sigma \)-representation: Consider \( X \sim (X, P_{XX})^n \). If \( P_{XX} > 0 \). Then a Min\( \sigma \)R of \( X \) is, for \( v := [v_1, \ldots, v_n]^T \in \mathbb{R}^n \) \( e_i \neq 0 \), the set \( \chi = \{ x_i, v_i \}_{i=1}^{n+1} \) with

\[ w_{n+1} = \frac{\sum_{i=1}^{n} (w_i + v_i)^2}{\sum_{i=1}^{n} v_i^2}, \quad \tilde{w} = w_{n+1} \left[ v_1^2, \ldots, v_n^2 \right]^T, \]

\[ \tilde{E} := \sqrt{P_{XX}} \left( I_n + v v^T \right)^{-\frac{1}{2}} \frac{\sqrt{W}}{2} \frac{1}{2} diag(v)^{-1}, \]

\[ e = -w_{n+1} \tilde{E} \tilde{w}, \quad [x_1, \ldots, x_{n+1}] = \{ E, e \} + [X]_{1 \times (n+1)}. \]

**Corollary 5** If \( w_i > 0 \), then the normalized Min\( \sigma \)R of Theorem 3 becomes

\[ w_{n+1} = \frac{1}{1 + \sum_{i=1}^{n} v_i^2}, \quad \tilde{w} = w_{n+1} \left[ v_1^2, \ldots, v_n^2 \right]^T, \]

\[ \tilde{E} := \sqrt{P_{XX}} \left( I_n + v v^T \right)^{-\frac{1}{2}} \frac{\sqrt{W}}{2} \frac{1}{2} diag(v)^{-1}, \]

\[ e = -w_{n+1} \tilde{E} \tilde{w}, \quad [x_1, \ldots, x_{n+1}] = \{ E, e \} + [X]_{1 \times (n+1)}. \]

If \( w_i > 0 \) and \( v = \rho C^{-1}[1]_{n \times 1} \), with \( \rho \) and \( C \) as in Tab I [6], then \( \chi \) is the minimum sigma set of [12].

The Mi\( \sigma \)R from Theorem 3 is currently the only consistent \( \sigma R \) constituted by less than \( 2n \) points, given that the set of [12] is a particular case of it and, to the best of the authors’ knowledge, the other reduced sets do not fit Definition 1, i.e. the mean and/or covariance matrix of their prior distribution are not matched (see Section II-D).
IV. UNSCENTED TRANSFORMATIONS

A. Unscented Transformation

In this section, we re introduce the definition of the Unscented Transformation (UT). In general terms, the UT consists of two sets of weighted points (the sigma points) that approximate a pdf of two rv's in the case where there is a functional dependence between them.

Henceforth, consider \( Y := f(X) \in \Phi^n, X \sim (\bar{X}, M_X^2, \ldots, M_X^l)^n \). For the sets \( \chi := \{ \chi_i, \psi_i(m) \psi_i(m)^T \} \), \( j \geq 2 \) \( N \) \( i = 1 \); \( \gamma := \{ \gamma_i, \psi_i(m) \psi_i(m)^T \} \gamma_i = f(\chi_i), j \geq 2 \) \( N \) \( i = 1 \); and vectors \( \lambda^\eta \in \{ \chi_1, \ldots, X_N, \gamma_1, \ldots, \gamma_N \}, \eta = 1, 2, \ldots \); the sample means are

\[
\mu_\chi := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) \chi_i, \quad \mu_\gamma := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) \gamma_i. \tag{19}
\]

For even \( j \), their sample central moments are

\[
M^j_\chi := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) (\chi_i - \mu_\chi)(\chi_i^T - \mu_\gamma)^{j/2}, \tag{20}
\]

\[
M^j_\gamma := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) (\gamma_i - \mu_\chi)(\gamma_i^T - \mu_\gamma)^{j/2}, \tag{21}
\]

\[
M^j_{\chi^1 \ldots \chi^j} := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) \otimes^j \left[ (\chi_i - \mu_\chi)(\chi_i^T - \mu_\gamma)^{1/2} \right]. \tag{22}
\]

For odd \( j \), they are

\[
M^j_\chi := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) (\chi_i - \mu_\chi)(\chi_i^T - \mu_\gamma)^{j+1/2} \otimes (\chi_i - \mu_\chi), \tag{23}
\]

\[
M^j_\gamma := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) (\gamma_i - \mu_\chi)(\gamma_i^T - \mu_\gamma)^{j+1/2} \otimes (\gamma_i - \mu_\gamma), \tag{24}
\]

\[
M^j_{\chi^1 \ldots \chi^j} := \frac{1}{N} \sum_{i=1}^{N} \psi_i(m) \otimes^j \left[ (\chi_i - \mu_\chi)(\chi_i^T - \mu_\gamma)^{1/2} \right]. \tag{25}
\]

Definition 2 (Unscented Transformation) Consider equations (19)-(25). If \( \mu_\chi = \bar{X} \) and \( M^2_\chi = M^2_\chi, j = 2, \ldots, l \), then the \( l \)th order Unscented Transformation (UT) is given by

\[
\text{UT}(f, \bar{X}, M^2_\chi, \ldots, M^l_\chi, \chi, \bar{X}, M^2_\chi, \ldots, M^l_\chi, \gamma, \bar{X}, M^2_\chi, \ldots, M^l_\chi, \eta) := [\mu_\gamma, M^2_\gamma, \ldots, M^l_\gamma, M^2_{\chi^2} \ldots, M^l_{\chi^l}, \chi, \bar{X}, M^2_{\chi^2} \ldots, M^l_{\chi^l}] .
\]

Remark 6 Every \( l \)th \( N \sigma R \) is a set \( \chi \) of a UT.

This form of defining the UT as a function mapping \( f, \bar{X}, P_X \) to the transformed sample mean and CM's can also be used in Monte Carlo and quadrature methods. Moreover, one should notice that negative weights can lead to negative values of the sample moments.

A UT can be viewed as a function that maps \( X \) and \( Y \) to sets which approximate the \((X, Y)\) joint pdf. For instance, a \( 2 \)UT can be viewed as the approximation (this interpretation is inspired on (46))

\[
\left( X \quad Y \right) \approx \left( \bar{X} \quad \bar{Y} \right) \sim \left( \left( \mu_\chi \right)^T, \Sigma_{\chi \chi}^T \Sigma_{\chi \gamma}^T \Sigma_{\gamma \gamma}^T \right)^T.
\]

The next theorem states the approximation quality for the \( Y \)'s. The notation \( Y \mid [\epsilon, f] \) stands for the Taylor Series of \( Y \) around \( X = \epsilon \) truncated at the \( l \)th order.

Theorem 4 If \( M^l_\chi = M^l_\chi, j = 2, \ldots, l, \mu_\chi = \bar{X} \) and \( f \) is \( l \)th differentiable, then:

\[
1) \quad \mu_{[\mu_\chi, f]} = \bar{Y}[X, \bar{X}],
\]

\[
2) \quad \Sigma_{[\mu_\chi, f]} = P_{XY}^{[X]} P_{XY}^{[X]} \quad \text{if } f \text{ is even},
\]

\[
3) \quad \Sigma_{[\mu_\chi, f]} = P_{XY}^{[X]} P_{XY}^{[X]} \quad \text{if } f \text{ is odd};
\]

\[
\mu_\chi = \bar{X} \quad \text{and} \quad M^p_\chi = M^p_\chi, p = 2, \ldots, l. \quad \text{For the first assertion,}
\]

\[
\mu_{[\mu_\chi, f]} = f(\mu_\chi) + \frac{1}{l+1} \sum_{i=1}^{l} \frac{2}{i!} \sum_{i_1 \ldots i_l=1}^{N} \left( M^2_{\chi^1 \ldots \chi^l} \right)_{i_1, \ldots, i_l} \frac{\partial^i f(x)}{\partial x^{i_1} \ldots \partial x^{i_l}} \bigg|_{x=\mu_\chi} + \ldots
\]

\[
+ \frac{1}{l+1} \sum_{i_1 \ldots i_l=1}^{N} \left( M^p_{\chi^1 \ldots \chi^l} \right)_{i_1, \ldots, i_l} \frac{\partial^i f(x)}{\partial x^{i_1} \ldots \partial x^{i_l}} \bigg|_{x=\mu_\chi}
\]

\[
\mu_{[\mu_\chi, f]} = \bar{Y}[X, \bar{X}],
\]

\[
\text{For the second, } \Sigma_{[\mu_\chi, f]} = \Theta_0^2 \Sigma_{\gamma \gamma} + \ldots + \Theta_2^2 \Sigma_{\gamma \gamma}, \text{ where } \Theta_0^2 \Sigma_{\gamma \gamma}
\]

\[
\text{is given in (27), shown at the top of the next.}
\]

For the third assertion, \( \Sigma_{\gamma \gamma}^{[X]} = \Theta_1^2 \Sigma_{\gamma \gamma} + \ldots + \Theta_2^2 \Sigma_{\gamma \gamma}, \text{ where } \]

\[
\Theta_0^2 \Sigma_{\gamma \gamma} = \frac{1}{l+1} \sum_{i_1 \ldots i_l=1}^{N} \left( M^2_{\chi^1 \ldots \chi^l} \right)_{i_1, \ldots, i_l} \frac{\partial^i f(x)}{\partial x^{i_1} \ldots \partial x^{i_l}} \bigg|_{x=\mu_\chi}.
\]

The remaining steps can be proven similarly.

Theorem 4 is the first to provide the estimation quality of the CCM, which is of the order \( l \) (\( \Sigma_{[\mu_\chi, f]} = P_{XY}^{[X]} \)).

The approximations of the posterior rv's are not guaranteed for any function \( f \), but only for the \( l \)th differentiable ones.

\[
\Theta_0^2 \Sigma_{\gamma \gamma} = \frac{1}{l+1} \sum_{i_1 \ldots i_l=1}^{N} \left( M^2_{\chi^1 \ldots \chi^l} \right)_{i_1, \ldots, i_l} \frac{\partial^i f(x)}{\partial x^{i_1} \ldots \partial x^{i_l}} \bigg|_{x=\mu_\chi}.
\]

(27)
For \( l = 2 \), the transformed CM is approximated up to order 1 \( (\Sigma_{\gamma}^{[\mu_2]} = P_{Y|X}^{[\mu_2]} ) \) (cf. Section II-C1).

According to Theorem 4, a sufficient condition for a second order approximation of the transformed CM is \( l = 4 \), since, for even \( l \), \( \Sigma_{\gamma}^{[\mu_2]} = P_{Y|X}^{[\mu_2]} \). In order to verify this, suppose \( \mu_\gamma = X \) and \( \Sigma_{\gamma}^{[\mu_2]} = M_\gamma \), \( i = 2, \ldots, 4 \). Then, from (27), \( \Sigma_{\gamma}^{[\mu_2]} = \Theta_{\Sigma_{\gamma}}^{[\mu_2]} + \cdots + \Theta_{\Sigma_{\gamma}}^{[\mu_2]} = 2n = P_{Y|X}^{[\mu_2]} \). Moreover, consider \( X \sim N(0, I_n, I_n) \), \( Y := f(X) = [x_1, \ldots, x_n]^T \) and suppose \( \mu_\gamma = X \) and \( \Sigma_{\gamma}^{[\mu_2]} = M_\gamma \), \( i = 2, \ldots, 6 \). Then, \( \Sigma_{\gamma}^{[\mu_2]} = \Theta_{\Sigma_{\gamma}}^{[\mu_2]} + \cdots + \Theta_{\Sigma_{\gamma}}^{[\mu_2]} = 15n = P_{Y|X}^{[\mu_2]} \).

This result does not imply that the mean and CM estimates of a 2UT are equal to the ones obtained through linearization. We can point out two reasons. First, for a 2UT, \( \mu_\gamma^{[\mu_2]} = \mu_\gamma^{[\mu_2]} \) for linearization. Second, even though both linearization and a 2UT have \( \Theta_{\Sigma_{\gamma}}^{[\mu_2]} = \Theta_{\Sigma_{\gamma}}^{[\mu_2]} \), it happens that, from (27), \( \Theta_{\Sigma_{\gamma}}^{[\mu_2]} \) and \( \Theta_{\Sigma_{\gamma}}^{[\mu_2]} \) are partially equal for a 2UT, but not for linearization (\( \Theta_{\Sigma_{\gamma}}^{[\mu_2]} = 0 \)).

B. Scaled Uncented Transformation

In this section, we redefine the Scaled Uncented Transformation (ScUT). This new definition is similar (not equal) to the AuxUT of [38] (Tab II [4:6]), since the SUT of [38] (Tab II [1:3]) cannot be applied to any previous sigma set (see Section II-E1). Furthermore, definitions similar to the previous SUT of [38] and to the UT of [35] (Tab I [7]) are presented at the end of this section as particular cases of the ScUT.

Unless otherwise specified, the term Scaled Uncented Transformation will henceforth refer to the following definition.

Definition 3 (Scaled Uncented Transformation) Consider, for \( \alpha \in [0, 1] \) and \( \kappa \in (0, 1] \), the function

\[
g(f, X, \beta, \alpha, \kappa) = \frac{f(\beta + \alpha(X - \beta)) - f(\beta)}{\kappa} + f(\beta),
\]

and the sets \( \chi := \{ \chi_1, \chi_2, \chi_3, \chi_4, X \} \) and \( \gamma := \{ \gamma_1, \gamma_2, \gamma_3, \gamma_4, \gamma_5, \gamma_6, \gamma_7, \gamma_8 \} \),

\[
\Sigma_{\Sigma_{\gamma}}^{[\alpha]} := \alpha \sum_{i=1}^{N} u_i \gamma_i (\chi_i - \mu_\gamma),
\]

\[
\Sigma_{\chi}^{[\alpha]} := \alpha \sum_{i=1}^{N} u_i \gamma_i (\chi_i - \mu_\gamma),
\]

where \( \Sigma_{\Sigma_{\gamma}}^{[\alpha]} \) is the scaled sample covariance matrix of \( \gamma \) and \( \Sigma_{\chi}^{[\alpha]} \) is the scaled sample cross covariance of \( \chi \) and \( \gamma \). If \( \mu_\gamma = X \) and \( \Sigma_{\chi}^{[\alpha]} = P_{Y|X} \), then the Scaled Uncented Transformation (ScUT) is defined by

\[
\text{ScUT}(f, X, P_{Y|X}) := [\mu_\gamma, \Sigma_{\Sigma_{\gamma}}^{[\alpha]}, \Sigma_{\chi}^{[\alpha]}].
\]

Remark 7 Every 2thNσR is a set \( \chi \) of a ScUT.

Such a definition for the CCM of the ScUT cannot be found for the scaled UT's of the literature. Crossing covariances are not treated in the SUT of [38] or in the AuxUT of [38]. In the UT of [35], the CCM is defined differently and restricted only to the symmetric set defined there (see Section II-E3).

In Section II-E2 and II-E3, it was shown that \( \alpha \) modifies the second order terms of both \( \Sigma_{\Sigma_{\gamma}}^{[\alpha]} \) and \( \Sigma_{\chi}^{[\alpha]} \). In order to check the influence of \( \alpha \) in the covariances of the ScUT, define \( \Theta_{\Sigma_{\gamma}}^{[\alpha]} \) and \( \Theta_{\Sigma_{\gamma}}^{[\alpha]} \) as the \( l \)th term of the Taylor Series of \( \Sigma_{\gamma}^{[\alpha]} \) and \( \Sigma_{\chi}^{[\alpha]} \), respectively, and consider the equations at the bottom of the page. The ScUT scales the terms of order 3 and higher for \( \mu_\gamma \) and of order 2 and higher for \( \Sigma_{\gamma}^{[\alpha]} \) and \( \Sigma_{\chi}^{[\alpha]} \). However, if \( \chi \) is symmetric, then \( \Sigma_{\gamma}^{[\alpha]} = [0, \cdots, 0] = \Theta_{\Sigma_{\gamma}}^{[\alpha]} = [0, \cdots, 0] \) and \( \alpha \) does not modify the second order of \( \Sigma_{\gamma}^{[\alpha]} \) (cf. Section II-E3).

The next theorem gives the estimation quality of the ScUT.

Theorem 5 If \( \Sigma_{\gamma}^{[\alpha]} = P_{X|X} \), \( \mu_\gamma = X \) and \( f \) is 2nd-order differentiable, then \( \mu_\gamma^{[\mu_2]} = P_{X|X}^{[\mu_2]} \), \( \Sigma_{\gamma}^{[\alpha]} = P_{X|X}^{[\mu_2]} \) and \( \Sigma_{\chi}^{[\alpha]} = P_{X|X}^{[\mu_2]} \). Furthermore, if \( X \) and \( \chi \) are symmetric, then \( \Sigma_{\gamma}^{[\alpha]} = P_{X|X}^{[\mu_2]} \).

Proof: Suppose \( \mu_\gamma = X \), \( \Sigma_{\chi}^{[\alpha]} = P_{X|X} \). For the first assertion, 

\[
\mu_\gamma = f(\mu_\gamma) + \sum_{i_1, i_2 = 1}^{n} \left( \frac{\partial f(x)}{\partial x_1} \right)_{x_1 = \mu_\gamma} \left( \frac{\partial f(x)}{\partial x_2} \right)_{x_2 = \mu_\gamma} + \cdots + \sum_{i_1, \cdots, i_l = 1}^{n} \left( \frac{\partial f(x)}{\partial x_1} \right)_{x_1 = \mu_\gamma} \cdots \left( \frac{\partial f(x)}{\partial x_l} \right)_{x_l = \mu_\gamma}
\]

\[
\Theta_{\gamma_{\Sigma_{\gamma}}}^{[\alpha]} = \sum_{i_1, i_2 = 1}^{n} \left( \frac{\partial f(x)}{\partial x_1} \right)_{x_1 = \mu_\gamma} \cdots \left( \frac{\partial f(x)}{\partial x_l} \right)_{x_l = \mu_\gamma}
\]

For the third assertion,

\[
\Sigma_{\chi}^{[\alpha]} = \sum_{i_1, \cdots, i_l = 1}^{n} \left( \frac{\partial f(x)}{\partial x_1} \right)_{x_1 = \mu_\gamma} \cdots \left( \frac{\partial f(x)}{\partial x_l} \right)_{x_l = \mu_\gamma}
\]

For the last assertion, note that \( X \) symmetric implies \( M_\chi^{[\alpha]} = [0, \cdots, 0] \) and \( \chi \) symmetric implies \( M_\chi^{[\alpha]} = [0, \cdots, 0] \)
Similar to the 2UT, the CM of the transformed rv is estimated only up to first order. Theorem 5 gives, for the first time, the estimation quality of the sample CCM. The next corollary states a new result.

**Corollary 6** A ScUT with sets \( \{x_i, w_i^a, w_i^c\}_{i=1}^{N+1} \) and \( \{y_i, w_i^a, w_i^c|y_i = g(f(x_i, \mu_x, \alpha)\}_{i=1}^{N+1} \) is a 2UT with sets \( \{x_i, w_i^a, w_i^c\}_{i=1}^{N+1} \) and \( \{y_i, w_i^a, w_i^c, w_i^{a,cc}\}_{i=1}^{N+1} \) where \( w_i^{a,cc} = \alpha w_i^a \) and \( w_i^{a,cc} = \alpha w_i^a \) are the weights to calculate the sample CM and CCM, respectively.

Because of the way these transformations were defined, every ScUT is a 2UT and, therefore, every result obtained for the 2UT can also be applied to the ScUT. We proceed by redefining the SUT of [38] and the UT of [35].

**Definition 4** Let \( \chi := (x_i, w_i)_{i=1}^{N} \) be a normalized \( \sigma R \) of \( X \) with \( \chi_N = \tilde{X} \), and consider the sets \( \chi^i := (x_i, w_i^a, w_i^c| x_i = \chi_N + \alpha (x_i - \chi_N)\}_{i=1}^{N} \) and \( \gamma^i := \{\gamma_i, w_i| \gamma_i = f(x_i)\}_{i=1}^{N} \) where, for \( \alpha \in (0, 1], w_N^i := \alpha^{-2}w_i + 1 - \alpha^{-2}, w_i = \alpha^{-2}w_i, i = 1, ..., N - 1 \). Define the modified sample CM and the sample CCM of \( \gamma^i \), respectively, as

\[
\Sigma_{\gamma^i} := \sum_{i=1}^{N} w_i (\gamma_i - \mu_{\gamma}) (\gamma_i - \mu_{\gamma})^T + (1 - \alpha^2) (\gamma_{N} - \mu_{\gamma}) (\gamma_{N} - \mu_{\gamma})^T
\]

and \( \Sigma_{\chi^i} := \sum_{i=1}^{N} w_i (\chi_i - \mu_{\chi})(\chi_i - \mu_{\chi})^T \). Then the Simplex Scated Unscented Transformation (SisScUT) is defined by

\[
\text{SisScUT}(f, X, P_{X\chi}, \alpha) := \{\gamma^i, \Sigma_{\gamma^i}, \Sigma_{\chi^i}\}.
\]

**Definition 5** Let \( \chi := (x_i, w_i)_{i=1}^{2n+1} \) with \( x_{2n+1} = \tilde{X} \) and \( w_{2n+1} = \lambda/(n+\lambda) \) for \( \alpha \in (0, 1], \lambda \in R \setminus \{0\} \) and \( \lambda = \alpha^2(n+\lambda) \). Consider the set \( \tilde{\gamma} := \{\gamma_i, w_i^m, w_i^c| \gamma_i = f(x_i)\}_{i=1}^{2n+1} \) where \( w_i^m := w_{2n+1}, w_i^c := w_{2n+1} + (1 - \alpha^{-2}), w_i^c := w_{2n+1} + (1 - \alpha^{-2}) \), and \( w_i^c \) is the weight for calculating the sample CCM. Then the Symmetric Intrinsically Scated Unscented Transformation (SisyScUT) is defined by

\[
\text{SisyScUT}(f, X, P_{X\chi}, \alpha) := \{\gamma_i, \Sigma_{\gamma^i}, \Sigma_{\chi^i}\}.
\]
Next, we introduce the Scaled SRUT and some results concerning this Transformation. This definition is necessary for the Scaled SRUKF’s (Section V-B), the first one in the literature.

**Definition 7 (Scaled Square-Root UT)** Consider the sets $\chi$ and $\gamma$ as in Definition 3 with $\mu_\chi = \bar{X}$ and $\Sigma_\chi = \sqrt{P_{XX}}\sqrt{P_{XX}}^T$. Given a matrix $\sqrt{Q}$, define $\Sigma_{\chi\gamma} := \Sigma_{\gamma\gamma} + \sqrt{Q}\sqrt{Q}^T$. Then the Scaled Square-Root Unscented Transformation (ScSRUT) is defined by

$$ \text{ScSRUT} \left( f, \bar{X}, \sqrt{P_{XX}}, \sqrt{Q}, \alpha \right) := \left[ \mu_\gamma, \sqrt{\Sigma_{\gamma\gamma}} \hat{S}_\gamma, \hat{S}_\gamma, \hat{S}_\gamma^T, \Sigma_\gamma \right]. $$

**Corollary 8** A ScSRUT with sets $\chi = \{ \chi_i, w_i^m, w_i^f \}_{i=1}^N$ and $\{ \gamma_i, w_i^m, w_i^f | \gamma_i = g(f(\chi_i, \mu_\chi, \alpha, \alpha^2))\}_{i=1}^1$ is a SRUT with the sets $\chi$ and $\{ \gamma_i, w_i^m, \alpha^2 w_i^f, \alpha w_i^f \}$.

**Remark 8** Every 2nd $N\sigma R$ is a set $\chi$ of a SRUT.

Finally, we state new ScSRUT results similar to the ones in Section IV-B for the particular transformations.

**Definition 8** Consider the sets $\chi'$ and $\gamma'$ as in Definition 4 with $\mu_{\chi'} = \bar{X}$ and $\Sigma_{\chi'\chi'} = \sqrt{P_{XX}}\sqrt{P_{XX}}^T$. Given a matrix $\sqrt{Q}$, define $\Sigma_{\gamma'\gamma'} := \Sigma_{\gamma'\gamma'} + \sqrt{Q}\sqrt{Q}^T$. Then the Simplex Scaled Square-Root Unscented Transformation (SiScSRUT) is defined by

$$ \text{SiScSRUT} \left( f, \bar{X}, \sqrt{P_{XX}}, \sqrt{Q}, \alpha \right) := \left[ \mu_{\gamma'}, \sqrt{\Sigma_{\gamma'\gamma'}} \hat{S}_{\gamma'}, \hat{S}_{\gamma'}, \hat{S}_{\gamma'}^T, \Sigma_{\gamma'\gamma'} \right]. $$

**Definition 9** Consider the sets $\hat{\chi}$ and $\hat{\gamma}$ as in Definition 5 with $\mu_{\hat{\chi}} = \bar{X}$ and $\Sigma_{\hat{\chi} \hat{\chi}} = \sqrt{P_{XX}}\sqrt{P_{XX}}^T$. Given a matrix $\sqrt{Q}$, define $\Sigma_{\hat{\gamma} \hat{\gamma}} := \Sigma_{\hat{\gamma} \hat{\gamma}} + \sqrt{Q}\sqrt{Q}^T$. Then the Symmetric Intrinsically Scaled Square-Root Unscented Transformation (SyInScSRUT) is defined by

$$ \text{SyInScSRUT} \left( f, \bar{X}, \sqrt{P_{XX}}, \sqrt{Q}, \alpha \right) := \left[ \mu_\gamma, \sqrt{\Sigma_{\gamma\gamma}} \hat{S}_\gamma, \hat{S}_\gamma, \hat{S}_\gamma^T, \Sigma_{\gamma\gamma} \right]. $$

**Corollary 9** Every SiScSRUT is an ScSRUT and every SyInScSRUT is an ScSRUT.

## V. UNSCENTED KALMAN FILTERS

In this section, we present discrete-time Unscented Kalman Filters. These are recursive applications of the UT (Section IV) to the discrete-time stochastic filtering problem (see Section II) in a KF framework.

### A. Unscented Kalman Filter

For the sake of clarity of presentation, we only explicitly describe the Unscented Kalman Filter for the system in additive form (1). The general filter for (2) can be similarly obtained by considering the augmented functions $f^n : \mathbb{R}^{n_x} \to \mathbb{R}^{n_x}$ and $h^n : \mathbb{R}^{n_y} \to \mathbb{R}^{n_y}$ such that, for $x_{k-1|k-1} := [\tilde{x}_{k-1|k-1}^T, \nu_{k-1|k-1}^T]^T$,

$$ f^n(x^n, k) = f(x^n_{(1),n_x}^k, x^n_{(n_x+1:n_x+n_u),1}^k, k), $$

$$ h^n(x^n, k) = h(x^n_{(1),n_x}^k, x^n_{(n_x+1:n_x+n_u),1}^k, k). $$

(29)

**Additive UKF (AdUKF):** Consider (1),

1) $\hat{x}_{k-1|k-1}$ and $\hat{P}_{xx}^{k-1|k-1}$ are given.

2) Obtain the predicted statistics of the state:

$$ \begin{cases} \hat{x}_{k|k-1} = \hat{P}_{xx}^{k|k-1} + Q_k. \\ \hat{P}_{xx}^{k|k-1} = \hat{P}_{xx}^{k-1|k-1}. \end{cases} $$

3) Obtain the predicted statistics of the measurement:

$$ \begin{cases} \tilde{y}_{k|k-1} = \tilde{y}_{k|k-1} + R_k. \\ \tilde{P}_{yy}^{k|k-1} = \tilde{P}_{yy}^{k-1|k-1} + R_k. \end{cases} $$

4) Obtain the corrected statistics of the state by (7).

**B. Square-Root Unscented Kalman Filter**

We now present the Square-Root Unscented Kalman Filter (SRUKF). The main difference between this filter and other types of UKF is the fact that the SRUKF propagate the square-root matrix of the CM’s directly, which is computationally more stable than squaring the propagated CM [48].

As Section II-F1 pointed out, the SRUKF’s in the literature present three steps in which Cholesky factors are downdated: in the calculations of the square-root matrices of the CM for the predicted state; in the CM for the innovation; and in the CM for the corrected state. While, in the first two cases, downdating is only performed when negative weights exist, the last one is always performed. Due to the fact that downdating steps can be computationally unstable (see Section II-F1), we derive an alternative form – which is an extension of the results of [48] and [6] – that uses the downdating procedure only for the negative weight components.

Define $S^+, S^-, S^+, S^-$ as in Section IV-C and note that

$$ \tilde{P}_{xx}^{k|k-1} = S^+ S^T - S^- S^T, \quad \tilde{P}_{yy}^{k|k-1} = S^\gamma S^\gamma T - S^- S^\gamma T $$

and $\tilde{P}_{xx}^{k|k-1} = S^+ S^T - S^- S^\gamma T + R_k$. Therefore,

$$ \tilde{P}_{xx}^{k|k} = [S^+ - G_k S^\gamma, G_k R][\bar{\sigma}]^T - [S^- - G_k S^\gamma, G_k R][\bar{\sigma}]^T, $$
which shows that $\hat{P}_{x|x}^{k|k}$ can be obtained through updating and downdating. The latter is only performed for the negative weight cases.

The SRUKF is presented below. It is more general than the algorithms currently in the literature, since these are restricted to the case where only the central weight, $w_0$, can be negative, whereas our SRUKF does not restrict the quantity of negative weights.

**Additive SRUKF (AdSRUKF):** Consider (1),

1. $\hat{x}_{k-1|k-1}$ and $\sqrt{P}_{x|x}^{k-1|k-1}$ are given.
2. Obtain the predicted statistics of the state:
   $$\left[\hat{x}_{k|k-1}, \sqrt{P}_{x|x}^{k|k-1}\right] = \text{SRUT}_1 \left(f, \hat{x}_{k-1|k-1}, \sqrt{P}_{x|x}^{k-1|k-1}, \sqrt{Q}_k\right).$$
3. Obtain the predicted statistics of the measurement:
   $$\left[\hat{y}_{k|k-1}, \sqrt{P}_{yy}^{k|k-1}, S_X^+, S_X^-, S_y^+, S_y^-, \hat{P}_{xy}^{k|k-1}\right] = \text{SRUT}_2 \left(h, \hat{x}_{k|k-1}, \sqrt{P}_{x|x}^{k|k-1}, \sqrt{T_k}\right).$$
4. Obtain the corrected statistics of the state by:
   $$G_k = \hat{P}_{xy}^{k|k-1} \sqrt{P}_{yy}^{k-1|k-1}^{-T} \sqrt{P}_{yy}^{k-1|k-1},$$
   $$\hat{x}_{k|k} = \hat{x}_{k|k-1} + G_k (y_k - \hat{y}_{k|k-1}),$$
   $$\sqrt{P}_{x|x}^{k|k} = \text{cu} \left(\left[S_X^+ - G_k S_y^+, \left[S_X^- - G_k S_y^-\right], G_k \sqrt{T_k}\right\right).$$

**C. Consistent Unscented Kalman Filter variations**

In order to clarify which UKF’s and SRUKF’s in the literature are consistent, we put all filter variants for $\text{UT}_1 = \text{UT}_2$ and $\text{SRUT}_1 = \text{SRUT}_2$ in Table IV. Each filter is the resulting variant of using the proposed AdUKF (AdSRUKF) with the corresponding UT (SRUT) in the first column and with the corresponding $\sigma$-representation in the first row (e.g. the Min. Sc. AdUKF, third row and fourth column, is the result of using the AdUKF with the ScUT and the Min(R)).

One should notice that consistent variants of the UKF (SRUKF) in the literature are particular cases of the proposed UKF (SRUKF) definitions in this work. Also, these definitions are able to provide new filter variants (e.g. the Scaled Square-Root Unscented Kalman Filters).

**D. Computational complexity and numerical implementations**

From the computational complexity point-of-view, the UKF’s most expensive operations are the square-root matrix operation of $P_{x|x}^{k-1|k-1} + Q_k [O(n^3)]$ and the matrix inversion of $P_{yy}^{k-1|k-1} [O(n y)]$, where $n_y$ is the dimension of the measurement vector). Hence, for the case in which $n_y \leq n$, the computational complexity of the UKF is $O(n^3)$, which is the same as for the EKF [43]. From a numerical implementation standpoint, even though the Cholesky decomposition seems to be the most adopted method to compute the square-root matrix of the CM for the state, some studies indicate that other methods, such as SVD decomposition, provide better estimation quality (see [61] for more details). Some code implementations are available on-line (e.g. [62] and [63]).

For the SRUKF, the computational complexity is also $O(n^3)$ due to the triangularization ($\text{tria}\{\}$), which is its most expensive operation. One example of triangularization is the QR decomposition, which has different implementations: the Householder QR requires $n^3/3$ floating points operations (flops) for a $n \times n$ matrix; the Givens QR requires $2n^3$ flops; and the modified Gram-Schmidt QR requires $2n^3$ flops [64]. However, in terms of numerical implementation, SRUKF’s behave better than the non-square root forms when implemented in a machine with poor precision [42].

**E. Higher-order Unscented Kalman Filters**

In this work, the AdUKF and AdSRUKF were defined only with 2nd order UT’s. Extensions to higher orders can be done in at least two ways. A first one is given by the following algorithm:

$l$th ($l > 2$) order (Gaussian) Additive UKF: Consider (1),

1. $\hat{x}_{k-1|k-1}$ and $\hat{P}_{x|x}^{k-1|k-1}$ are given.
2. Choose $l \in N$, $l > 2$ and obtain the central moments $\hat{M}_{x|k-1|k-1}^2$, ..., $\hat{M}_{x|k-1|k-1}^l$ for $x_{k-1|k-1} \sim N(\hat{x}_{k-1|k-1}, \hat{P}_{x|x}^{k-1|k-1} + Q_k)$.
3. Obtain the predicted statistics of the state:
   $$\left[\hat{x}_{k|k-1}, \hat{P}_{x|x}^{k|k-1}\right] = \text{IUT}_1 \left(f, \hat{x}_{k-1|k-1}, \hat{M}_{x|k-1|k-1}^2, ..., \hat{M}_{x|k-1|k-1}^l\right).$$
4. Obtain the central moments $\hat{M}_{x|k-1|k-1}^2$, ..., $\hat{M}_{x|k-1|k-1}^l$ for $x_{k|k-1} \sim N(\hat{x}_{k|k-1}, \hat{P}_{x|x}^{k|k-1} + R_k)$.
5. Obtain the predicted statistics of the measurement:
   $$\left[\hat{y}_{k|k-1}, \hat{P}_{xy}^{k|k-1}\right] = \text{IUT}_2 \left(h, \hat{x}_{k|k-1}, \hat{M}_{x|k-1|k-1}^2, ..., \hat{M}_{x|k-1|k-1}^l\right).$$
6. Obtain the corrected statistics of the state by (7).

This approach uses the Gaussian assumption of the Kalman Filter to obtain the previous first $l$ moments of the state for each $\text{IUT}$. Generally, higher values of $l$ result in a larger number of sigma-points and better state estimation (cf. Theorem 4). Note that the higher-order UKF of [41] is a particular case of this proposed filter for the scalar case.

A second way is to propagate, at every time step, not only the mean and the covariance matrix of the state, but also its higher-order moments up to the chosen $l$th order (a similar approach that does not use UT’s is proposed by [65]). This method does not assume that the state follows a Gaussian distribution at every time step and provides a better approximation when compared to the first one, but at the cost of increased effort in developing the recursive equations and also of having a computationally more expensive algorithm.

**VI. CONCLUSION**

From the study of the state-of-the-art in Unscented estimation theory, we were able to observe several important problems (Sections II-B to II-F) concerning (I) the matching order of the transformed covariance matrix (CM) (Sections II-C1 and II-E2) and the transformed cross-covariance matrix (CCM) (Sections II-C2 and II-E3) of both the Unscented
The authors would like to thank Mr. Pedro Henrique R. Q. Santana for proof-reading the paper.

REFERENCES


Henrique M. T. Menegaz received the B.S. degree in electrical engineering and the M.S. degree in engineering of electronic systems and automation from the University of Brasilia (UnB), Brazil, in 2007 and 2011, respectively, where he is currently pursuing the Ph.D. degree.

His major field of study is filtering of nonlinear dynamic systems and robotics.

João Y. Ishihara received the Ph.D. degree in Electrical Engineering from the University of São Paulo, Brazil, in 1998.

He is currently an Associate Professor at the University of Brasília, Brazil. His research interests include robust filtering and control theory, singular systems, and robotics.

Geovany A. Borges received the Ph.D. degree from the Universidade de Brasilia, Brazil. His main research interests are robotics, stochastic filtering, and estimation.

Alessandro N. Vargas received Ph.D. degree in electrical engineering from the University of Campinas, Brazil, in 2009. He is currently an Associate Professor at the Universidade Tecnológica Federal do Paraná, Brazil. His research interests include stochastic systems control and their applications.